

Growth deletion models for the Web graph and other massive networks

Changping Wang

Technical Report CS-2005-03

April, 2005

Faculty of Computer Science 6050 University Ave., Halifax, Nova Scotia, B3H 1W5, Canada

Growth deletion models for the web graph and other massive networks

April 12, 2005

Changping Wang¹

Department of Mathematics & Statistics, Dalhousie University

Halifax, NS, Canada, B3H 3J5

Abstract: We propose new evolutionary stochastic models for the web graph and other massive networks, where edges are deleted over time and an edge is chosen to be deleted with probability inversely proportional to the in-degree of the destination. The degree distributions of graphs generated by our models follow a power law. A rigorous proof of power law degree distributions is given using martingales and concentration results. Depending on the parameters, the exponent of the power law can be any number in $(1, \infty)$. For this reason, our models apply not only to the web graph, but to certain biological networks, where the power law exponent is in the interval (1, 2).

Keywords. web graph, power law graphs, degree distributions, scale free networks, stochastic graph models

1 Introduction

In the past few years, there has been much interest in understanding the properties of real-world large-scale networks such as the structure of the Internet and the World Wide Web. It has been observed that many such networks have a so-called power law degree distribution: the proportion of nodes of degree k is approximately $\frac{1}{k\gamma}$, where $\gamma > 1$ is a fixed real number. Such graphs are sometimes called scale-free in the literature. A graph is called a power law graph if the fraction of nodes with degree k is proportional to $\frac{1}{k\gamma}$ for some constant $\gamma > 0$. The standard models of random graphs introduced by Erdős and Rényi [11] and Gilbert [12] are not appropriate for studying these networks, since they generate graphs which, with high probability, have binomial degree distributions.

¹This research was partially supported by MITACS as part of the MoMiNIS project.

A large number of power law random graph models [1, 3, 5, 10, 13] have been proposed. For two recent surveys on models of the web graph, see [4, 6]. In all of these models, at each time step nodes and edges are added, but never *deleted*. An evolving graph model incorporating in its design both the addition and deletion of nodes and edges may more accurately model the evolution of the web graph. Recently, the models of [7, 9] incorporate the addition and deletion of nodes *during* the generation of nodes. We refer to such models as *growth-deletion models*.

In [7], Chung, Lu introduced a growth-deletion model $G(p_1, p_2, p_3, p_4, m)$, for undirected graphs with parameters m a positive integer, and probabilities p_1 , p_2 , p_3 , p_4 satisfying $p_1 + p_2 + p_3 + p_4 = 1$, $p_3 < p_1$, and $p_4 < p_2$. The graph H is a fixed nonempty graph. To form G_{t+1} , they proceed as follows. With probability p_1 , add a new node v_{t+1} and m edges from v_{t+1} to existing nodes chosen with probability proportional to their degrees. We refer to this as preferential attachment. With probability p_2 , add m new edges with endpoints to be chosen among existing nodes by preferential attachment. With probability p_3 , delete a node chosen uniformly at random (u.a.r). With probability p_4 , delete m edges chosen u.a.r.

In [9], Cooper et al. introduced an undirected model with three parameters α , α_0 and α_1 which generates a sequence of simple graphs G_t , t = 1, 2, ..., where the graph $G_t = (V_t, E_t)$ has v_t nodes and e_t edges. They start with G_1 consisting of an isolated node x_1 . At time t, with probability $1 - \alpha - \alpha_0$, they delete a randomly chosen node x_1 from V_{t-1} . If $V_{t-1} = \emptyset$, they do nothing; with probability α_0 , they delete min $\{m, |E_{t-1}|\}$ randomly chosen edges from E_{t-1} ; with probability α_1 , they add a node x_t with m random edges incident with x_t to G_{t-1} , the endpoints are chosen by preferential attachment; with probability $\alpha - \alpha_1$, they add m random edges to existing nodes. The endpoints are chosen by preferential attachment. The models of [7, 9] generate scale-free graphs whose exponent γ is in the interval $[2, \infty)$.

However, we observe that in realistic networks such as the web graph, new nodes are more likely to join existing nodes with high degree, while links pointing to a node with high degree are less likely to be deleted. Motivated by this observation, we propose a new directed model, called a biased edge-deletion model, where we delete a directed edge with probability inversely proportional to the in-degree of the destination. Hence, edges pointing to "popular" nodes (that is, nodes with high in-degree) are less likely to be deleted.

We describe these network models precisely in Section 2, and state our main results there. In Section 3, a rigorous proof for the power law degree distributions is given using martingales and concentration results. We emphasize that our models generate graphs with power law exponent in the interval $(1, \infty)$. This is significant since certain massive networks, such as the network of protein-protein interaction networks in a living cell, have power law exponents the interval (1, 2); see [8]. Hence, our models are used not only as models of the web graph, but for many other massive networks.

2 Edge-deletion models

In this section, we introduce three edge-deletion models and state our main result. Assuming that α and β are two nonnegative real numbers satisfying $\alpha + \beta < 1$ and $\beta < \frac{1}{2}$, we consider a random process which generates a sequence of graphs G_t , t = 0, 1, 2, ..., ... The graph $G_t = (V_t, E_t)$ will have n_t nodes and e_t edges.

Model 1: To initialize the process, at t = 0 we start with an initial digraph G_0 with n_0 nodes and m_0 edges.

At time t, with probability $1 - \alpha - \beta$ we add a node v_t to G_{t-1} , with a directed loop. With probability α we add a directed edge uv to the existing nodes, where the origin is chosen with probability proportional to its out-degree and the destination is chosen proportional to its in-degree. With probability β , if $e_{t-1} > 0$, we delete a directed edge, where an edge is chosen inversely proportional to the in-degree of the destination; if $e_{t-1} = 0$, we do nothing.

Model 2: This model is defined similarly to Model 1 except that edges to be deleted u.a.r.

The next model generates undirected graphs.

Model 3: To initialize the process, at t = 0 we start with an initial graph G_0 with n_0 nodes and m_0 edges.

At time t, with probability $1-\alpha-\beta$ we add a node v_t to G_{t-1} along with an edge. An endpoint is v_t , the other endpoint is chosen by preferential attachment. With probability α we add an edge uv to the existing nodes. The endpoints u and v are chosen by preferential attachment. With probability β , if $e_{t-1} > 0$, we delete an edge u.a.r; if $e_{t-1} = 0$, we do nothing.

Note: Model 3 is the same as those models in [7] with $p_3 = 0$ and m = 1, and in [9] with $\alpha + \alpha_0 = 1$ and m = 1. We obtain the same result as those in [7, 9].

Denote by $d_{k,t}^{in}$ the number of nodes with in-degree k at time t in Model 1 and Model 2, and denote by $d_{k,t}$ the number of nodes with degree k at time t in Model 3. In the following theorem, we will show that, asymptotically, $d_{k,t}^{in}$ and $d_{k,t}$ follow a power law. If A is an event in a probability space, then we write Pr(A) for the probability of A in the space.

Theorem 1 For the models 1, 2, and 3, we have the following.

1. For Model 1, the in-degree distribution follows a power law with exponent $\gamma = 1 + \frac{1-2\beta}{\alpha} \in (1,\infty)$. More precisely, we have

$$Pr\left(\left|d_{k,t}^{in} - b_{k,1}' n_t\right| > 2\epsilon \sqrt{t}(1 + b_{k,1}')\right) < 4e^{-\epsilon^2/2},$$

where

$$b'_{k,1} = \frac{b_{k,1}}{1 - \alpha - \beta} = (1 + O(k^{-1})) \frac{C_1(\alpha, \beta)}{1 - \alpha - \beta} k^{-\gamma},$$

and $C_1(\alpha, \beta)$ is a constant.

2. For Model 2, the in-degree distribution follows a power law with exponent $\gamma = 1 + \frac{1-2\beta}{\alpha-\beta} \in (1,\infty)$. More precisely,

$$Pr\left(\left|d_{k,t}^{in} - b_{k,2}' n_t\right| > 2\epsilon\sqrt{t}(1 + b_{k,2}')\right) < 4e^{-\epsilon^2/2}$$

where

$$b'_{k,2} = \frac{b_{k,2}}{1 - \alpha - \beta} = (1 + O(k^{-1})) \frac{C_2(\alpha, \beta)}{1 - \alpha - \beta} k^{-\gamma},$$

and $C_2(\alpha, \beta)$ is a constant.

3. For Model 3, the degree distribution follows a power law with exponent $\gamma = 1 + \frac{2-4\beta}{1+\alpha-3\beta} \in (1,\infty)$. More precisely,

$$Pr\left(\left|d_{k,t} - b'_{k,3}n_t\right| > 2\epsilon\sqrt{t}(1 + b'_{k,3})\right) < 2(e^{-\epsilon^2/2} + e^{-\epsilon^2/8}),$$

where

$$b'_{k,3} = \frac{b_{k,3}}{1 - \alpha - \beta} = (1 + O(k^{-1})) \frac{C_3(\alpha, \beta)}{1 - \alpha - \beta} k^{-\gamma},$$

and $C_3(\alpha, \beta)$ is a constant.

3 Proof of Theorem 1

As t increases, G_t may be defined recursively. For each t, let τ_t be a random variable of G_t . Let c be a positive integer. The random variable τ_t is said to satisfy the c-Lipschitz condition if

$$|\tau_{t+1}(G_{t+1}) - \tau_t(G_t)| \le c$$

whenever G_{t+1} is obtained from G_t by adding or deleting some edges or some nodes at time t+1. The proof of Theorem 1 will follow by the next Lemma, which is the Azuma-Hoeffding Inequality. See for example, Theorem 7.2.1 of [2]. If X is a random variable, then we denote E(X) for its expected value.

Lemma 2 If τ_t satisfies the c-Lipschitz condition, then for every $\delta > 0$

$$Pr[|\tau_t - E(\tau_t)| > \delta\sqrt{t}] < 2e^{-\frac{\delta^2}{2c^2}}.$$

In particular, τ_t is almost surely very close to its expected value $E(\tau_t)$ with an error term $o(t^{\frac{1}{2}+\delta})$ for any $\delta > 0$, as t approaches infinity.

The next two lemmas are useful in our proof of Theorem 1, and also serve as a warm up for the application of Lemma 2.

Lemma 3 For each of the models 1, 2, and 3, we have the following.

1. For $t \geq 0$, the expected value of the number of (directed)edges e_t at time t is

$$E(e_t) = m_0 + (1 - 2\beta)t.$$

2. For every $\epsilon > 0$,

$$Pr[|e_t - E(e_t)| > \epsilon t^{\frac{2}{3}}] < 2e^{-\frac{\epsilon^2}{2}t^{\frac{1}{3}}}.$$

Proof Define

$$X_j = \begin{cases} 1 & \text{an edge is added at time } j; \\ -1 & \text{an edge is deleted at time } j. \end{cases}$$

So, $e_t = m_0 + \sum_{j=1}^t X_j$. Therefore, $E(e_t) = m_0 + \sum_{j=1}^t E(X_j)$. We know that with probability β we delete an edge and with probability $1 - \alpha - \beta + \alpha = 1 - \beta$ we add an edge. So, $E(X_j) = 1 - 2\beta$ for each j. Hence, (1) holds. Since e_t satisfies the 1-Lipschitz condition, by Lemma 2 with $\delta = \epsilon t^{1/6}$, (2) holds.

Lemma 4 For each of the models 1, 2, and 3, we have the following.

1. For $t \geq 0$, the expected number of nodes n_t at time t is

$$E(n_t) = n_0 + (1 - \alpha - \beta)t.$$

2. For every $\epsilon > 0$,

$$Pr(|n_t - E(n_t)| > \epsilon \sqrt{t}) < 2e^{-\frac{\epsilon^2}{2}}.$$

Proof Define

$$X_j = \begin{cases} 1 & \text{a node is added at time } j; \\ 0 & \text{otherwise.} \end{cases}$$

So, $n_t = n_0 + \sum_{j=1}^t X_j$. Therefore, $E(n_t) = n_0 + \sum_{j=1}^t E(X_j)$. We know that with probability $1 - \alpha - \beta$ we add a new node. So, $E(X_j) = 1 - \alpha - \beta$ for each j. Hence, (1) holds. Since n_t satisfies the 1-Lipschitz condition, by Lemma 2 with $\delta = \epsilon$,(2) holds.

Proof of Theorem 1. We prove (1) first. For the sequence of random variables $\{d_{k,t}^{in}\}$, we will compute the corresponding expected value $E(d_{k,t}^{in})$ here. At time 0, there is an initial graph G_0 with n_0 nodes and m_0 edges. Let $d_{k,0}^{in} = d_k^0$ be the number of nodes with in-degree $k, k \geq 0$ at time 0. At time 1, a node with a loop is added. We abbreviate "with probability" by "w.p.". Assume that there are e_t edges at time t, for $t \geq 0$. It is not hard to see that

$$d_{0,t+1}^{in} = \begin{cases} d_{0,t}^{in} + 1 & \text{w.p. } \beta \frac{d_{1,t}^{in}}{e_t}; \\ d_{0,t}^{in} & \text{otherwise.} \end{cases}$$

and

$$d_{1,t+1}^{in} = \begin{cases} d_{1,t}^{in} + 1 & \text{w.p. } 1 - \alpha - \beta + \beta \frac{d_{2,t}^{in}}{e_t}; \\ d_{1,t}^{in} - 1 & \text{w.p. } (\alpha + \beta) \frac{d_{1,t}^{in}}{e_t}; \\ d_{1,t}^{in} & \text{otherwise.} \end{cases}$$

In general, for k > 1, $d_{k,t+1}^{in}$ can increase by 1 because a node of in-degree k-1 receives an edge or a node of in-degree k+1 loses an edge; $d_{k,t+1}^{in}$ can decrease by 1 because a node of in-degree k receives an edge or loses an edge. Thus we have that

$$d_{k,t+1}^{in} = \begin{cases} d_{k,t}^{in} + 1 & \text{w.p. } \alpha \frac{(k-1)d_{k-1,t}^{in}}{e_t} + \beta \frac{d_{k+1,t}^{in}}{e_t}; \\ d_{k,t}^{in} - 1 & \text{w.p. } \alpha \frac{kd_{k,t}^{in}}{e_t} + \beta \frac{d_{k,t}^{in}}{e_t}; \\ d_{k,t}^{in} & \text{otherwise.} \end{cases}$$

Hence,

$$E(d_{k,t+1}^{in}|G_t) = d_{k,t}^{in} \left(1 - \frac{k\alpha + \beta}{e_t}\right) + \alpha \frac{(k-1)d_{k-1,t}^{in}}{e_t} + \beta \frac{d_{k+1,t}^{in}}{e_t}.$$
 (3.1)

Define $\overline{e}_t = m_0 + (1 - 2\beta)t$ and let $A_t = \{|e_t - \overline{e}_t| \le \epsilon t^{\frac{2}{3}}\}$ be an event. Define

$$a_t = \begin{cases} 1 & A_t \text{ occurs;} \\ 0 & \text{otherwise.} \end{cases}$$

By Lemma 3, we know that

$$Pr(a_t = 1) \ge 1 - 2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}.$$
 (3.2)

By (3.1), we obtain

$$d_{k,t}^{in} \left(1 - \frac{k\alpha + \beta}{\overline{e}_t - \epsilon t^{\frac{2}{3}}} \right) + \frac{\alpha(k-1)d_{k-1,t}^{in} + \beta d_{k+1,t}^{in}}{\overline{e}_t + \epsilon t^{\frac{2}{3}}} \le E(d_{k,t+1}^{in}|G_t, a_t = 1)$$

$$\leq d_{k,t}^{in} \left(1 - \frac{k\alpha + \beta}{\overline{e}_t + \epsilon t^{\frac{2}{3}}} \right) + \frac{\alpha(k-1)d_{k-1,t}^{in} + \beta d_{k+1,t}^{in}}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}. \tag{3.3}$$

It is easy to see that

$$E(d_{k,t+1}^{in}|G_t) = P(a_t = 1)E(d_{k,t+1}^{in}|G_t, a_t = 1) + P(a_t = 0)E(d_{k,t+1}^{in}|G_t, a_t = 0).$$
(3.4)

Note that $d_{k,t}^{in} - 1 \le E(d_{k,t+1}^{in}|G_t) \le d_{k,t}^{in} + 1$. So,

$$E(d_{k,t}^{in}) - 1 \le E(d_{k,t+1}^{in}) \le E(d_{k,t}^{in}) + 1. \tag{3.5}$$

Taking expectation on both sides of (3.4), together with (3.2), (3.3) and (3.5), we obtain

$$\left(E(d_{k,t}^{in})\left(1 - \frac{k\alpha + \beta}{\overline{e_t - \epsilon t^{\frac{2}{3}}}}\right) + \frac{\alpha(k-1)E(d_{k-1,t}^{in}) + \beta E(d_{k+1,t}^{in})}{\overline{e_t + \epsilon t^{\frac{2}{3}}}}\right)\left(1 - 2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right)$$

$$\leq E(d_{k,t+1}^{in}) \leq (E(d_{k,t}^{in}) + 1) \left(2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right) +$$

$$E(d_{k,t}^{in})\left(1 - \frac{k\alpha + \beta}{\overline{e}_t + \epsilon t^{\frac{2}{3}}}\right) + \frac{\alpha(k-1)E(d_{k-1,t}^{in}) + \beta E(d_{k+1,t}^{in})}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}.$$

$$(3.6)$$

Let $E(d_{k,t}^{in}) = b_{k,1}t + c_{k,t}$, where $c_{k,t} = o(t)$ is a lower order term. To choose an appropriate value for $b_{k,1}$, we substitute it into (3.6) and let t tend to infinity. We obtain, for k > 1,

$$\beta b_{k+1,1} - (1 - \beta + k\alpha)b_{k,1} + (k-1)\alpha b_{k-1,1} = 0.$$
(3.7)

In the following we will solve (3.7) by using the Laplace Method. This method was first used in the study of web graph models by [9]. Replacing k by k + 1 in (3.7), we get

$$\beta b_{k+2,1} + [-\alpha(k+1) + \beta - 1]b_{k+1,1} + k\alpha b_{k,1} = 0,$$

which is of the form

$$(A_2(k+2) + B_2)b_{k+2,1} + (A_1(k+1) + B_1)b_{k+1,1} + (A_0k + B_0)b_{k,1} = 0. (3.8)$$

where $A_2 = 0, B_2 = \beta, A_1 = -\alpha, B_1 = \beta - 1, A_0 = \alpha, B_0 = 0$. We make the substitution

$$b_{k,1} = \int_{a}^{b} t^{k-1} v(t) dt, \tag{3.9}$$

where a, b are constants, and v(t) is a function of t to be determined. Integrating by parts, we obtain

$$kb_{k,1} = [t^k v(t)]_a^b - \int_a^b t^k v'(t)dt.$$
(3.10)

Let $\phi_1(t) = A_2 t^2 + A_1 t + A_0$ and $\phi_0(t) = B_2 t^2 + B_1 t + B_0$. Substituting (3.9) and (3.10) into (3.8), we can get

$$[t^{k}\phi_{1}(t)v(t)]_{a}^{b} - \int_{a}^{b} t^{k}\phi_{1}(t)v'(t)dt + \int_{a}^{b} t^{k-1}\phi_{0}(t)v(t)dt = 0.$$
(3.11)

If we ensure that

$$\frac{v'(t)}{v(t)} = \frac{\phi_0(t)}{t\phi_1(t)},\tag{3.12}$$

and

$$[t^k v(t)\phi_1(t)]_a^b = 0, (3.13)$$

then (3.8) will be satisfied. Now (3.13) can be satisfied by choosing a=0 and b equal to a root of $v(t)\phi_1(t)=0$. Moreover, since $A_2=0, B_2=\beta, A_1=-\alpha, B_1=\beta-1, A_0=\alpha, B_0=0$, we can obtain

$$\phi_1(t) = A_2 t^2 + A_1 t + A_0 = (t - 1)(-\alpha)$$

and

$$\phi_0(t) = B_2 t^2 + B_1 t + B_0 = \beta t^2 + (\beta - 1)t.$$

Thus, we have the following differential equation.

$$\frac{v'(t)}{v(t)} = \frac{\phi_0(t)}{t\phi_1(t)} = \frac{\beta t + \beta - 1}{\alpha(1 - t)}.$$
(3.14)

Integrating (3.14), we obtain

$$v(t) = Ce^{-\frac{\beta}{\alpha}t}(1-t)^{\gamma}.$$

where $\gamma=\frac{1-2\beta}{\alpha}$ and C is a constant. For convenience, we choose C=1. With this choice of v(t), we can choose b=1 and (3.13) is satisfied. So, we have a=0,b=1 and $v(t)=e^{-\frac{\beta}{\alpha}t}(1-t)^{\gamma}$.

Now we go back to (3.9) and determine b_k as follows.

$$b_{k,1} = \int_0^1 t^{k-1} v(t) dt$$

$$= \int_0^1 t^{k-1} e^{-\frac{\beta}{\alpha}t} (1-t)^{\gamma} dt$$

$$= \int_0^1 t^{k-1} (1-t)^{\gamma} \sum_{j=0}^{\infty} \frac{(-\frac{\beta}{\alpha}t)^j}{j!} dt$$

$$= \sum_{j=0}^{\infty} \frac{(-\frac{\beta}{\alpha})^j}{j!} \int_0^1 t^{k+j-1} (1-t)^{\gamma} dt$$

$$= \sum_{j=0}^{\infty} \frac{(-\frac{\beta}{\alpha})^j}{j!} \frac{\Gamma(k+j)\Gamma(\gamma+1)}{\Gamma(k+j+\gamma+1)}$$

$$= \sum_{j=0}^{\infty} \frac{(-\frac{\beta}{\alpha})^j \Gamma(\gamma+1)}{j!} \frac{\Gamma(k+j)}{\Gamma(k+j+\gamma+1)}.$$

Let $\{a_n\}$ and $\{b_n\}$ be two sequences of real numbers, we write $a_n \approx b_n$ if

$$\lim_{n \to \infty} \frac{a_n}{b_n} = 1.$$

Assuming that k is large, then we can use Stirling's formula for $\Gamma(k+j)$ and $\Gamma(k+j+\gamma+1)$ as follows.

$$\Gamma(k+j) \approx \sqrt{2\pi}(k+j-1)^{k+j-\frac{1}{2}}e^{-(k+j-1)}$$

and

$$\Gamma(k+j+\gamma+1) \approx \sqrt{2\pi}(k+j+\gamma)^{k+j+\gamma+\frac{1}{2}}e^{-(k+j+\gamma)}$$

Hence, we obtain

$$b_{k,1} = \sum_{j=0}^{\infty} \frac{(-\frac{\beta}{\alpha})^{j} \Gamma(\gamma+1)}{j!} \frac{\Gamma(k+j)}{\Gamma(k+j+\gamma+1)}$$

$$= (1+O(k^{-1})) \sum_{j=0}^{\infty} \frac{e^{1+\gamma}(-\frac{\beta}{\alpha})^{j} \Gamma(\gamma+1)}{j!} (k+\gamma+j)^{-\gamma-1}$$

$$= (1+O(k^{-1})) C_{1}(\alpha,\beta) k^{-\gamma-1},$$

where $C_1(\alpha, \beta)$ is a constant.

The sequence of random variables $\{d_{k,t}^{in}\}$ satisfies the 1-Lipschitz condition. By Lemma 2, for every $\epsilon > 0$, we obtain

$$Pr(|d_{k,t}^{in} - b_{k,1}t| > 2\epsilon\sqrt{t}) < 2e^{-\epsilon^2/2},$$
 (3.15)

By Lemma 4, for every $\epsilon > 0$, we have

$$Pr(|n_t - n_0 - (1 - \alpha - \beta)t| > \epsilon \sqrt{t}) < 2e^{-\frac{\epsilon^2}{2}}.$$

With t large enough so that $n_0 \le \epsilon \sqrt{t}$, thus we obtain

$$Pr(|n_t - (1 - \alpha - \beta)t| > 2\epsilon\sqrt{t}) < 2e^{-\frac{\epsilon^2}{2}}.$$
 (3.16)

By (3.15) and (3.16), we obtain

$$Pr\left(\left|d_{k,t}^{in} - \frac{b_{k,1}}{1 - \alpha - \beta}n_t\right| > 2\epsilon\sqrt{t}\left(1 + \frac{b_{k,1}}{1 - \alpha - \beta}\right)\right) < 4e^{-\epsilon^2/2},$$

The proof of (1) follows.

Now we prove (2). For the sequence of random variables $\{d_{k,t}^{in}\}$, we will compute the corresponding expected value $E(d_{k,t}^{in})$ here. At time 0, there is an initial graph G_0 with n_0 nodes and m_0 edges. Let $d_{k,0}^{in} = d_k^0$ be the number of nodes with in-degree $k, k \geq 0$ at time 0. At time 1, a node with a loop is added. Assume that there are e_t edges at time t, for $t \geq 0$. It is not hard to see that

$$d_{0,t+1}^{in} = \begin{cases} d_{0,t}^{in} + 1 & \text{w.p. } \beta \frac{d_{1,t}^{in}}{e_t}; \\ d_{0,t}^{in} & \text{otherwise.} \end{cases}$$

and

$$d_{1,t+1}^{in} = \begin{cases} d_{1,t}^{in} + 1 & \text{w.p. } 1 - \alpha - \beta + \beta \frac{2d_{2,t}^{in}}{e_t}; \\ d_{1,t}^{in} - 1 & \text{w.p. } (\alpha + \beta) \frac{d_{1,t}^{in}}{e_t}; \\ d_{1,t}^{in} & \text{otherwise.} \end{cases}$$

In general, for k > 1, we have that

$$d_{k,t+1}^{in} = \begin{cases} d_{k,t}^{in} + 1 & \text{w.p. } \alpha \frac{(k-1)d_{k-1,t}^{in}}{e_t} + \beta \frac{(k+1)d_{k+1,t}^{in}}{e_t}; \\ d_{k,t}^{in} - 1 & \text{w.p. } (\alpha + \beta) \frac{kd_{k,t}^{in}}{e_t}; \\ d_{k,t}^{in} & \text{otherwise.} \end{cases}$$

Hence,

$$E(d_{k,t+1}^{in}|G_t) = d_{k,t}^{in} \left(1 - \frac{k(\alpha + \beta)}{e_t}\right) + \alpha \frac{(k-1)d_{k-1,t}^{in}}{e_t} + \beta \frac{(k+1)d_{k+1,t}^{in}}{e_t}.$$
 (3.17)

Let A_t and a_t be as defined in case (1). By (3.17), we obtain

$$d_{k,t}^{in}\left(1 - \frac{k(\alpha + \beta)}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}\right) + \frac{\alpha(k-1)d_{k-1,t}^{in} + \beta(k+1)d_{k+1,t}^{in}}{\overline{e}_t + \epsilon t^{\frac{2}{3}}} \le E(d_{k,t+1}^{in}|G_t, a_t = 1)$$

$$\leq d_{k,t}^{in} \left(1 - \frac{k(\alpha + \beta)}{\overline{e}_t + \epsilon t^{\frac{2}{3}}} \right) + \frac{\alpha(k-1)d_{k-1,t}^{in} + \beta(k+1)d_{k+1,t}^{in}}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}. \tag{3.18}$$

Using (3.4), together with (3.2), (3.18) and (3.5), we obtain

$$\left(E(d_{k,t}^{in})\left(1 - \frac{k(\alpha + \beta)}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}\right) + \frac{\alpha(k-1)E(d_{k-1,t}^{in}) + \beta(k+1)E(d_{k+1,t}^{in})}{\overline{e}_t + \epsilon t^{\frac{2}{3}}}\right)\left(1 - 2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right)$$

$$\leq E(d_{k,t+1}^{in}) \leq (E(d_{k,t}^{in}) + 1) \left(2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right) +$$

$$E(d_{k,t}^{in})\left(1 - \frac{k(\alpha + \beta)}{\overline{e}_t + \epsilon t^{\frac{2}{3}}}\right) + \frac{\alpha(k-1)E(d_{k-1,t}^{in}) + \beta(k+1)E(d_{k+1,t}^{in})}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}.$$
 (3.19)

Let $E(d_{k,t}^{in}) = b_{k,2}t + c_{k,t}$, where $c_{k,t} = o(t)$ is a lower order term. To choose an appropriate value for $b_{k,2}$, we substitute it into (3.19) and let t tend to infinity. We obtain, for k > 1,

$$(k+1)\beta b_{k+1,2} - (1-2\beta + k(\alpha+\beta))b_{k,2} + (k-1)\alpha b_{k-1,2} = 0.$$
(3.20)

Solving (3.20) by the Laplace Method, we obtain

$$b_{k,2} = (1 + O(k^{-1}))C_2(\alpha, \beta)k^{-\gamma_2 - 1},$$

where $C_2(\alpha, \beta)$ and $\gamma_2 = \frac{1-2\beta}{\alpha-\beta}$ are constants.

The sequence of random variables $\{d_{k,t}^{in}\}$ satisfies the 1-Lipschitz condition. By Lemma 2, for every $\epsilon > 0$, we obtain

$$Pr(|d_{k,t}^{in} - b_{k,2}t| > 2\epsilon\sqrt{t}) < 2e^{-\epsilon^2/2},$$
 (3.21)

By (3.21) and (3.16), we obtain

$$Pr\left(\left|d_{k,t}^{in} - \frac{b_{k,2}}{1 - \alpha - \beta}n_t\right| > 2\epsilon\sqrt{t}\left(1 + \frac{b_{k,2}}{1 - \alpha - \beta}\right)\right) < 4e^{-\epsilon^2/2},$$

The proof of (2) follows.

Finally we prove (3). For the sequence of random variables $\{d_{k,t}\}$, we will compute the corresponding expected value $E(d_{k,t})$ here. At time 0, there is an initial graph G_0 with n_0 nodes and m_0 edges. Let $d_{k,0} = d_k^0$ be the number of nodes with degree $k, k \geq 0$ at time 0. At time 1, a node with an edge is added. Assume that there are e_t edges at time t, for $t \geq 0$. It is not hard to see that

$$d_{0,t+1} = \begin{cases} d_{0,t} + 1 & \text{w.p. } \beta \frac{d_{1,t}}{e_t}; \\ d_{0,t} & \text{otherwise.} \end{cases}$$

and

$$d_{1,t+1} = \begin{cases} d_{1,t} + 2 & \text{w.p. } \beta \frac{N_2}{e_t}; \\ d_{1,t} + 1 & \text{w.p. } (1 - \alpha - \beta)(1 - \frac{d_{1,t}}{e_t}) + \beta \frac{2d_{2,t}}{e_t}; \\ d_{1,t} - 1 & \text{w.p. } \alpha(1 - \frac{d_{1,t}}{e_t}) + \beta \frac{d_{1,t} - 2N_1 - N_3}{e_t}; \\ d_{1,t} - 2 & \text{w.p. } \alpha(\frac{d_{1,t}}{2e_t})^2 + \beta \frac{N_1}{e_t}; \\ d_{1,t} & \text{otherwise.} \end{cases}$$

where N_1 , N_2 and N_3 denote the number of edges between nodes with degree 1, the number of edges between nodes with degree 2 and the number of edges between nodes with degree 2 and 1 at time t, respectively.

In general, for k > 1, $d_{k,t+1}$ can increase by 2 if we delete an edge whose two endpoints have degree k+1 or we add an edge whose two endpoints have degree k-1; $d_{k,t+1}$ can increase by 1 if we delete an edge whose exactly one endpoint has degree k+1 or we add an edge whose exactly one endpoint has degree k-1; $d_{k,t+1}$ can decrease by 2 if we delete or add an edge whose two endpoints have degree k; $d_{k,t+1}$ can decrease by 1 if we delete or add an edge whose exactly one endpoint has degree k; otherwise, $d_{k,t+1}$ stays at the same. Thus we have that

$$d_{k,t} + 2 \quad w.p. \quad \alpha \frac{((k-1)d_{k-1,t})^2}{4(e_t)^2} + \beta \frac{K_1}{e_t};$$

$$d_{k,t} + 1 \quad w.p. \quad \alpha \frac{(k-1)d_{k-1,t}}{e_t} \left(1 - \frac{kd_{k,t} + (k-1)d_{k-1,t}}{2e_t}\right) + \beta \frac{(k+1)d_{k+1,t} - 2K_1 - K_3}{e_t}$$

$$+ \frac{(1-\alpha-\beta)(k-1)d_{k-1,t}}{2e_t};$$

$$d_{k,t} - 1 \quad w.p. \quad \alpha \frac{kd_{k,t}}{e_t} \left(1 - \frac{kd_{k,t} + (k-1)d_{k-1,t}}{2e_t}\right) + \beta \frac{kd_{k,t} - 2K_2 - K_3}{e_t}$$

$$+ \left(1 - \alpha - \beta\right) \frac{kd_{k,t}}{2e_t};$$

$$d_{k,t} - 2 \quad w.p. \quad \alpha \frac{(kd_{k,t})^2}{4(e_t)^2} + \beta \frac{K_2}{e_t};$$

$$d_{k,t} \quad otherwise.$$

where K_1 , K_2 and K_3 denote the number of edges between nodes with degree k + 1, the number of edges between nodes with degree k and the number of edges between nodes with degree k + 1 and k at time t, respectively. Hence,

$$E(d_{k,t+1}|G_t) = d_{k,t} \left(1 - \frac{k(1+\alpha+\beta)}{2e_t} \right) + \frac{(1+\alpha-\beta)(k-1)d_{k-1,t} + 2\beta(k+1)d_{k+1,t}}{2e_t}.$$
(3.22)

By (3.22), we obtain

$$d_{k,t}\left(1 - \frac{k(1+\alpha+\beta)}{2(\overline{e}_t - \epsilon t^{\frac{2}{3}})}\right) + \frac{(1+\alpha-\beta)(k-1)d_{k-1,t} + 2\beta(k+1)d_{k+1,t}}{2(\overline{e}_t + \epsilon t^{\frac{2}{3}})} \le E(d_{k,t+1}|G_t, a_t = 1)$$

$$\leq d_{k,t} \left(1 - \frac{k(1+\alpha+\beta)}{2(\overline{e}_t + \epsilon t^{\frac{2}{3}})} \right) + \frac{(1+\alpha-\beta)(k-1)d_{k-1,t} + 2\beta(k+1)d_{k+1,t}}{2(\overline{e}_t - \epsilon t^{\frac{2}{3}})}.$$
(3.23)

Note that $d_{k,t} - 2 \le E(d_{k,t+1}|G_t) \le d_{k,t} + 2$. So,

$$E(d_{k,t}) - 2 \le E(d_{k,t+1}) \le E(d_{k,t}) + 2. \tag{3.24}$$

Note that (3.4) is also true for the sequence of random variables $\{d_{k,t}\}$. Using (3.4), together with (3.2), (3.23) and (3.24), we obtain

$$\left(E(d_{k,t})\left(1 - \frac{k(1+\alpha+\beta)}{2(\overline{e}_t - \epsilon t^{\frac{2}{3}})}\right) + \frac{(1+\alpha-\beta)(k-1)E(d_{k-1,t}) + 2\beta(k+1)E(d_{k+1,t})}{2(\overline{e}_t + \epsilon t^{\frac{2}{3}})}\right)\left(1 - 2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right)$$

$$\leq E(d_{k,t+1}) \leq (E(d_{k,t}) + 2) \left(2e^{-\frac{\epsilon^2 t^{\frac{1}{3}}}{2}}\right) +$$

$$E(d_{k,t})\left(1 - \frac{k(1+\alpha+\beta)}{2(\overline{e}_t + \epsilon t^{\frac{2}{3}})}\right) + \frac{(1+\alpha-\beta)(k-1)E(d_{k-1,t}) + 2\beta(k+1)E(d_{k+1,t})}{\overline{e}_t - \epsilon t^{\frac{2}{3}}}.$$
(3.25)

Let $E(d_{k,t}) = b_{k,3}t + c_{k,t}$, where $c_{k,t} = o(t)$ is a lower order term. To choose an appropriate value for $b_{k,3}$, we substitute it into (3.25) and let t tend to infinity. We obtain, for k > 1,

$$2\beta(k+1)b_{k+1,3} - (2-4\beta + (1+\alpha+\beta)k)b_{k,3} + (k-1)(1+\alpha-\beta)b_{k-1,3} = 0.$$
(3.26)

Solving (3.26) by the Laplace Method, we obtain

$$b_{k,3} = (1 + O(k^{-1}))C_3(\alpha, \beta)k^{-\gamma_3 - 1},$$

where $C_3(\alpha, \beta)$ and $\gamma_3 = \frac{2-4\beta}{1+\alpha-3\beta}$ are constants.

The sequence of random variables $\{d_{k,t}\}$ satisfies the 2-Lipschitz condition. By Lemma 2, for every $\epsilon > 0$, we obtain

$$Pr(|d_{k,t} - b_{k,3}t| > 2\epsilon\sqrt{t}) < 2e^{-\epsilon^2/8},$$
 (3.27)

By (3.27) and (3.16), we obtain

$$Pr\left(\left|d_{k,t} - \frac{b_{k,3}}{1 - \alpha - \beta}n_t\right| > 2\epsilon\sqrt{t}\left(1 + \frac{b_{k,3}}{1 - \alpha - \beta}\right)\right) < 2(e^{-\epsilon^2/2} + e^{-\epsilon^2/8}),$$

The proof of (3) follows.

4 Conclusion and discussion

In this paper, we use techniques in random graph theory to analyze power law graphs, and we solve recurrences by the Laplace method. Our models generate graphs with power law exponent in the interval $(1, \infty)$. This is a larger interval than the interval $[2, \infty)$ found for the models in [7, 9]. This is significant since certain massive networks, such as the network of protein-protein interaction networks in a living cell, have power law exponents the interval (1, 2); see [8]. Hence, our models can be used not only as models of the web graph, but for many other massive networks.

A problem that we cannot presently solve is how to rigorously analyze models that incorporate the deletion of nodes over time. For instance, one might want to make the deletion of high degree nodes less likely than low degree nodes. We expect to pursue this problem in future work.

Acknowledgement. I am grateful to Jeannette Janssen and Anthony Bonato for their constant guidance and encouragement.

References

- W. Aiello, F. Chung, L. Lu, Random evolution in massive graphs, Handbook on Massive Data Sets, (Eds. James Abello et al.), Kluwer Academic Publishers, (2002), 97-122.
- [2] N. Alon, J. H. Spencer, The Probabilistic Method, Wiley and Sons, New York, 1992.
- [3] A. Barabási, R. Albert, Emergence of scaling in random networks, *Science*, **286** (1999), 509-512.
- [4] B. Bollobás, O. Riordan, Mathematical results on scale-free graphs, *Handbook of graphs and networks* (S. Bornholdt, H. Schuster eds.), Wiley-VCH, Berlin (2002).

- [5] B. Bollobás, O. Riordan, J. Spencer, G. Tusnády, The degree sequence of a scale-free random graph process, *Random Structures and Algorithms*, **18** (2001), 270-290.
- [6] A. Bonato, A survey of models of the web graph, *Proceedings of Combinatorial and Algorithmic Aspects of Networking*, 2004.
- [7] F. Chung, L. Lu, Coupling on-line and off-line analyses for random power law graphs, submitted.
- [8] F. Chung, L. Lu, T. G. Dewey, D. J. Galas, Duplication models for biological networks, *Journal of Computational Biology*, **10** (2003), 677-688.
- [9] C. Cooper, A. Frieze, J. Vera, Random deletion in a scale free random graph process, submitted.
- [10] C. Cooper, A. Frieze, A general model of undirected web graphs, *Random Structures* and *Algorithms*, **22** (2003), 311-335.
- [11] P. Erdős, A. Rényi, On random graphs I., Publicationes Mathematicae Debrecen, 5 (1959), 290-297.
- [12] E.N. Gilbert, Random graphs, Annals of Mathematical Statistics, **30** (1959), 1141-1144.
- [13] R. Kumar, P. Raghavan, S. Rajagopalan, D. Sivakumar, A. Tomkins, E. Upfal, Stochastic models for the Web graph, *Proceedings of the 41st Annual Symposium on Foundations of Computer Science* (FOCS2003).